RSS Journals Club, September 30th 2013 A likelihood-based sensitivity analysis for publication bias in meta-analysis Applied Statistics 2013, 62, 47-62 John Copas University of Warwick, UK jbc@stats.warwick.ac.uk

## The paper

- discusses the statistical difficulties of publication bias, essentially a problem of non-random sampling
- suggests a sensitivity analysis based on a sample selection model
- 1. Introduction
- 2. Example a classic meta-analysis debacle (§2)
- 3. Selection models for publication bias  $(\S2-4)$
- 4. Example revisited  $(\S5.1)$
- 5. Discussion  $(\S5.3 \text{ and } 6)$

The three stages of meta analysis:

- Literature search and systematic review of relevant studies
- Statistical summary of each study
  - Study estimates  $\hat{\theta}_i$
  - Within-study variances  $\sigma_i^2$
- Combining summary statistics into an overall inference

– fixed effects model

$$\hat{\theta}_i \sim N(\theta, \sigma_i^2)$$

$$- \text{MLE} = \tilde{\theta} = \frac{\sum w_i \hat{\theta}_i}{\sum w_i}, \quad w_i = \frac{1}{\sigma_i^2}$$
$$- \text{Var}\{\tilde{\theta}\} = \frac{1}{\sum w_i}$$

Example: Yusuf et al. (1993)

Meta-analysis of 15 clinical trials on the effectiveness of intravenous magnesium in acute myocardial infarction

$$\theta = \log \frac{P(\text{death} \mid \text{treatment})}{P(\text{death} \mid \text{control})}$$

Relative risk = 
$$\exp{\{\tilde{\theta}\}} = .58(.46, .73)$$
  
P-value  $\approx 2 \times 10^{-6}$ 

Published conclusion: "magnesium is an effective, safe, simple and inexpensive intervention that should be introduced into clinical practice without delay" But then .....

ISIS-4 (1995), a very large multi-centre randomized clinical trial, reported mortality rates

- 2216/29011 (magnesium)
- 2103/29039 (control)
- Relative risk = 1.06(0.99, 1.13)
- P-value  $\approx 0.09$

Conclusion: there is **no significant difference**, magnesium may in fact be harmful.



To appear in a meta analysis a study has to be

- written up
- submitted
- accepted for publication
- found by the reviewer

## Conjecture

Studies reporting a significant result are more likely to survive this selection process

 $\Rightarrow$  the meta analysis will be biased

## Selection model for publication bias

- There is a population of studies  $(\hat{\theta}, \sigma^2)$  from which the *n* observed studies are a (non-random) selection
- The probability that a study is selected may depend on its t-statistic  $y = \hat{\theta}/\sigma$

 $\Rightarrow$  P(selected | study with  $\hat{\theta}, \sigma^2) = a(y)$ 

for some function a(y)

Examples

a(y) = 1 (no bias) $a(y) = \begin{cases} 1 & \text{if } y \le k \\ 0 & \text{if } y > k \end{cases} \text{ (negative bias)}$ 

Under the null hypothesis  $H_0: \theta = 0$ , for each study

$$y = \frac{\hat{\theta}}{\sigma} \sim N(0, 1)$$

Then under  $H_0$ 

$$P(\text{selection}) = p = \int a(y)\phi(y)dy$$
$$E(y|\text{selection}) = \mu = \frac{\int ya(y)\phi(y)dy}{p}$$

So under  $H_0$ 

 $E(\hat{\theta}|study selected) = \mu\sigma$ 



Probit random effects selection model

$$\hat{\theta} | \sigma \sim N(\theta, \sigma^2 + \tau^2)$$
  
 $P(\text{select} | \hat{\theta}, \sigma) = \Phi(\alpha + \beta \hat{\theta} / \sigma)$ 

Then the paper shows that

• 
$$p \approx \left(\frac{1}{n} \sum \left[\Phi\left\{\frac{\alpha + \beta\theta/\sigma_i}{\{1 + \beta^2(1 + \tau^2/\sigma_i^2)\}^{\frac{1}{2}}}\right\}\right]^{-1}\right)^{-1}$$

• Log likelihood is

$$L = -\frac{1}{2} \sum \log(\tau^2 + \sigma_i^2) - \frac{1}{2} \sum \frac{(\hat{\theta}_i - \theta)^2}{\tau^2 + \sigma_i^2}$$
$$+ \sum \log \Phi(\alpha + \beta \hat{\theta}_i / \sigma_i) - \sum \log \Phi \left\{ \frac{\alpha + \beta \theta / \sigma_i}{\{1 + \beta^2 (1 + \tau^2 / \sigma_i^2)\}^{\frac{1}{2}}} \right\}$$

**Statistical difficulty**: the available data (funnel plot) usually gives very little information about the value of p (the overall proportion of studies which are selected).

Sensitivity analysis. Fix the value of p and find the corresponding maximum likelihood estimates of the other parameters. Then

- Plot the confidence interval for  $\theta$  against p.
- Superimpose the fitted values  $E(\hat{\theta}|select, \sigma, p)$  on the funnel plot for a selection of values of p

#### Sensitivity analysis

For any given value of p we can get

- MLE  $\hat{\theta}_p$
- Confidence limits  $\{\hat{\theta}_p^{(L)}, \hat{\theta}_p^{(U)}\}$  based on  $2\{\max L_p L_p(\theta)\} \sim \chi_1^2$
- Fitted values: estimate of  $E(\hat{\theta}|select, \sigma, p) =$

$$\theta + \beta \sigma \frac{1 + \tau^2 / \sigma^2}{\sqrt{1 + \beta^2 (1 + \tau^2 / \sigma^2)}} \lambda \left( \frac{\alpha + \beta \theta / \sigma}{\sqrt{1 + \beta^2 (1 + \tau^2 / \sigma^2)}} \right)$$

 $(\lambda = \text{Mills ratio } \phi/\Phi)$ 





# Discussion — sensitivity analysis versus bias correction

- why don't we estimate all the parameters and hence find the MLE of p?
- why don't we use one of the selection models in the literature to find the MLE of  $\theta$ ?





### General comments

- many meta analyses suffer from publication bias, but this is almost always ignored
- publication bias usually means that the treatment effect is exaggerated
- it is impossible to adjust for publication bias unless we make un-testable assumptions
- 'selection by significance'  $\Rightarrow a(y)$
- the sensitivity analysis conditions on an interpretable parameter
- sensitivity analyses tend to be more robust to modelling assumptions than bias correction methods

## In the magnesium example

- standard meta-analysis gives a strongly significant (and strongly misleading) result
- there is evidence of a 'small study effect': smaller studies tend to give stronger treatment effects than the one reasonably larger study
- the selection model explains the funnel plot trend (e.g. fitted values of  $\theta$  when p = 0.5 all lie within the individual confidence intervals)
- the treatment effect is no longer significant if p < 0.6
- the sensitivity analysis suggests that the evidence remains significant provided there are less than about 9 missing studies, but at a much more modest level of significance

### References

ISIS-4 Collaborative Group (1995) A randomized factorial trial assessing early oral captopril, oral mononitrate and intravenous magnesium sulphate in 58,050 patients with suspected myocardial infarction. *Lancet*, **345**, 669-685.

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Yusuf S, Koon T, Woods K (1993) Intravenuos magnesium in acute myocardial infarction: an effective, safe, simple and inexpensive intervention. *Circulation*, **87**, 2043-2046.